



## Association Between the ERCC1 Polymorphisms and Glioma Risk: A Meta-Analysis of Case-Control Studies

LIU Yan<sup>1#</sup>, CAI Xiao Qin<sup>2#</sup>, ZHAO Lian Ying<sup>3</sup>, SHEN Heng Shan<sup>2</sup>, HU Jian Wei<sup>4\*</sup>

<sup>1</sup>Department of Epidemiology, School of Public Health, Medical College of Soochow University, Suzhou, Jiangsu, China

<sup>2</sup>Department of Diagnostic Center, the Kunshan Affiliated Hospital of Nanjing University of Traditional Chinese Medicine, Suzhou 215300, Jiangsu, China

<sup>3</sup>Health Supervision Institute of Kunshan, Kunshan, Jiangsu, China

<sup>4</sup>Maternal and Child Health Bureau of Kunshan, Kunshan, Jiangsu, China

#These authors contributed equally to this work and should be considered as co-first authors.

\*Corresponding author: HU Jian Wei, Physician, Maternal and Child Health Bureau of Kunshan, Kunshan, Jiangsu, China, Tel: 086-512-5733-6557; E-mail: [hujianwei19850826@163.com](mailto:hujianwei19850826@163.com)

### Abstract

**Objective:** Polymorphisms in DNA repair genes have been shown to influence DNA repair processes and to modify cancer susceptibility. Published data regarding the association between excision repair cross-complementing rodent repair deficiency complementation group1 (ERCC1) polymorphisms and glioma risk have been inconsistent and inconclusive. To acquire a more precise effect of the association between these polymorphisms and glioma risk, a meta-analysis was performed.

**Methods:** Data was collected in PubMed and EMBASE, with the last search up to 30<sup>th</sup> August 2013. A total of 6 studies were identified with 2642 cases and 3669 controls for ERCC1 C8092A polymorphism, and 4 studies were identified with 1390 cases and 1546 controls for ERCC1 C118T polymorphism. All of the statistical analyses were performed using statistical data 10.0.

**Results:** The combined results showed that ERCC1 C8092A polymorphism was associated with glioma risk (additive model: OR = 1.10, 95% CI 1.02 - 1.20; recessive model: OR = 1.51, 95% CI 1.24 - 1.85; co-dominant model AA vs. CC: OR = 1.52, 95% CI 1.24 - 1.86). As for ethnicity subgroup analysis, ERCC1 C8092A polymorphism was associated with increased glioma risk among Chinese (additive model: OR = 1.15, 95% CI 1.01 - 1.30; recessive model: OR = 1.34, 95% CI 1.02 - 1.75; AA vs. CC: OR = 1.37, 95% CI 1.03 - 1.81), and so among Caucasian except additive model (recessive model: OR = 1.75, 95% CI 1.31 - 2.34; co-dominant model AA vs. CC: OR = 1.70, 95% CI 1.27 - 2.29). No evidence of an association of ERCC1 C118T polymorphism with glioma was found.

**Conclusion:** The meta-analysis suggested that ERCC1 C8092A polymorphism might be associated with the increased risk of glioma, whereas ERCC1 C118T polymorphism might have no influence on the susceptibility of glioma.

### Keywords

Polymorphism, ERCC1, Glioma, Meta-analysis

### Introduction

Glioma is the most common and fatal neurological cancer [1,2] and accounts for almost 80% of primary malignant brain tumors [3,4]. Despite many advances in surgical and medical therapy in recent years, glioma consistently remains a fatal disease and there is no significant increase in the survival for patients with glioma [1,2]. Though several risk factors have been found, the exact pathogenesis of glioma remains unclear [5]. However, there is no doubt that genetic factors play important roles in the development of glioma [6,7]. Numerous studies are addressing associations of polymorphisms in DNA repair genes and cancer risk [8] because accurate and efficient DNA repair is crucial to genomic integrity and fidelity. The DNA repair system is complex, governed by more than 125 genes, many of which are polymorphic [9-12]. The DNA repair gene, excision repair cross-complementing rodent repair deficiency complementation group1 (ERCC1), whose products are important in the process of nucleotide excision repair (NER) lies on chromosome 19q13.3 in a putative glioma suppression region [13]. ERCC1 gene is reported to be a crucial gene in the NER pathway, and ERCC1 polymorphisms can modify the function of NER pathway, thus influence the risk of human cancers [14]. ERCC1 C8092A and C118T are two common polymorphisms of ERCC1 gene and they alters risk of several types of cancers [15-18]. Associations between polymorphisms in ERCC1 gene and glioma risk have also been examined but the observed associations were inconsistent [19-22], and a single study may be too underpowered to detect a possible small effect of the polymorphisms on glioma, especially when the sample size was relatively small. To the best of our knowledge, there has been no comprehensive quantitative assessment of the association of ERCC1 C8092A or C118T polymorphism with glioma risk. Hence, we performed a meta-analysis of all eligible studies to derive a more precise estimation of the associations of ERCC1 C8092A or C118T polymorphism with glioma risk.

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## Materials and Methods

### Literature search

PubMed and EMBASE were searched (the last search update on the 30<sup>th</sup> August 2013) using the search terms: 'ERCC1' and 'glioma or brain tumor'. All studies matching the eligible criteria were retrieved, and bibliographies checked for other relevant publications. Review articles and bibliographies of other relevant studies identified were hand searched to identify additional studies. Only published studies with full text were included in this meta-analysis.

### Inclusion criteria

The inclusion criteria were (a) evaluation of the ERCC1 C8092A or C118T polymorphism and glioma risk, (b) case-control study design, (c) human subject studies and (d) the size of the sample, odds ratios (ORs) and their 95% confidence intervals (CIs) or the information that can help infer the results in the papers.

### Data extraction

Information was carefully extracted from all eligible publications independently by two investigators (Jun Liu and Ting Lai), according to the inclusion criteria listed above. The following data was collected from each study: first author's surname, publication date, country, ethnicity, study design, genotyping methods, total number of cases and controls, and numbers of cases and controls with ERCC1 C8092A and C118T genotypes, respectively. For those studies that included subjects of different ethnic groups, data were extracted separately for each of the ethnic groups, categorized as Caucasians or Chinese. We did not define any minimum number of patients for including a study in our meta-analysis. When the same patient population was included in several studies only the most recent or complete study was included in this meta-analysis.

### Statistical analysis

Crude ORs with 95% CIs were calculated, according to the method of Woolf [23] to assess the association of ERCC1 C8092A and C118T polymorphisms with glioma risk. The pooled ORs for ERCC1 C8092A were performed for additive model (A allele vs. C allele), co-dominant model (AA vs. CC, AC vs. CC), dominant model (AA/AC vs. CC), recessive model (AA vs. AC/CC). As for ERCC1 C118T all these estimates were also calculated. Heterogeneity assumption was checked by a chi-square-based Q-test [24]. A p-value of more than 0.10 for the Q-test indicated a lack of heterogeneity across the studies, so the pooled estimation of the ORs of each study was calculated by the fixed effects model (Mantel-Haenszel method). Otherwise, the random effects model

(DerSimonian and Laird method) was used [25]. The significance of the pooled OR was determined by the Z-test, and  $P < 0.05$  was considered as statistically significant. To evaluate the ethnic-specific effect, subgroup analysis was conducted on the basis of different ethnicities. One-way sensitivity analysis was performed to assess the stability of the results, namely, a single study in the meta-analysis was deleted each time to reflect the influence of the individual data set to the pooled OR [26,27]. An estimate of the potential publication bias was carried out by funnel plot, in which the standard error (SE) of log (OR) of each study was plotted against its log (OR). An asymmetric plot suggested a possible publication bias. The funnel plot asymmetry was assessed by Egger's test, a linear regression approach to measure funnel plot asymmetry on the natural logarithmic scale of the OR. The significance of the intercept was determined by the t-test suggested by Egger,  $P < 0.05$  was considered representative of statistically significant publication bias [28]. Hardy-Weinberg equilibrium (HWE) in the control group was tested by the chi-square test for goodness of fit, and a p-value of  $< 0.05$  was considered significant. All the statistical tests for our meta-analysis were performed with STATA version 10.0 (Stata Corporation, College Station, TX). All p-values were two-sided.

## Results

### Study characteristics

Based on the search criterion, 84 articles were found. Through the step of screening the articles (Figure 1), seven articles [20-22,29-32] were included. Characteristics of included studies are summarized in table 1. Among of these studies, a total of 6 studies involving 2642 cases and 3669 controls for ERCC1 C8092A, while 4 studies involving of 1390 cases and 1546 controls for ERCC1 C118T were analyzed. For ERCC1 C8092A, 3 studies were carried on Chinese population, and other 3 studies were carried on Caucasian population. For ERCC1 C118T, 4 studies were provided data on Chinese population. MassARRAY was used to validate genotype in two papers, and TaqMan was used to validate genotype in two studies. Mixed PCR method was used to validate genotype in two studies and PCR-SSCP was used in one study. Genotype distributions in the control populations were all in agreement with HWE.

### Meta-analysis results

The result of this meta-analysis is shown in table 2. For ERCC1 C8092A polymorphism, Q-test was used in all of the genetic models and there was no significant heterogeneity. Therefore, the pooled ORs were calculated using fixed-effects model. No association for ERCC1 C8092A polymorphism and glioma risk was found in these genetic

**Table 1:** Main characteristics of all studies included in the meta-analysis

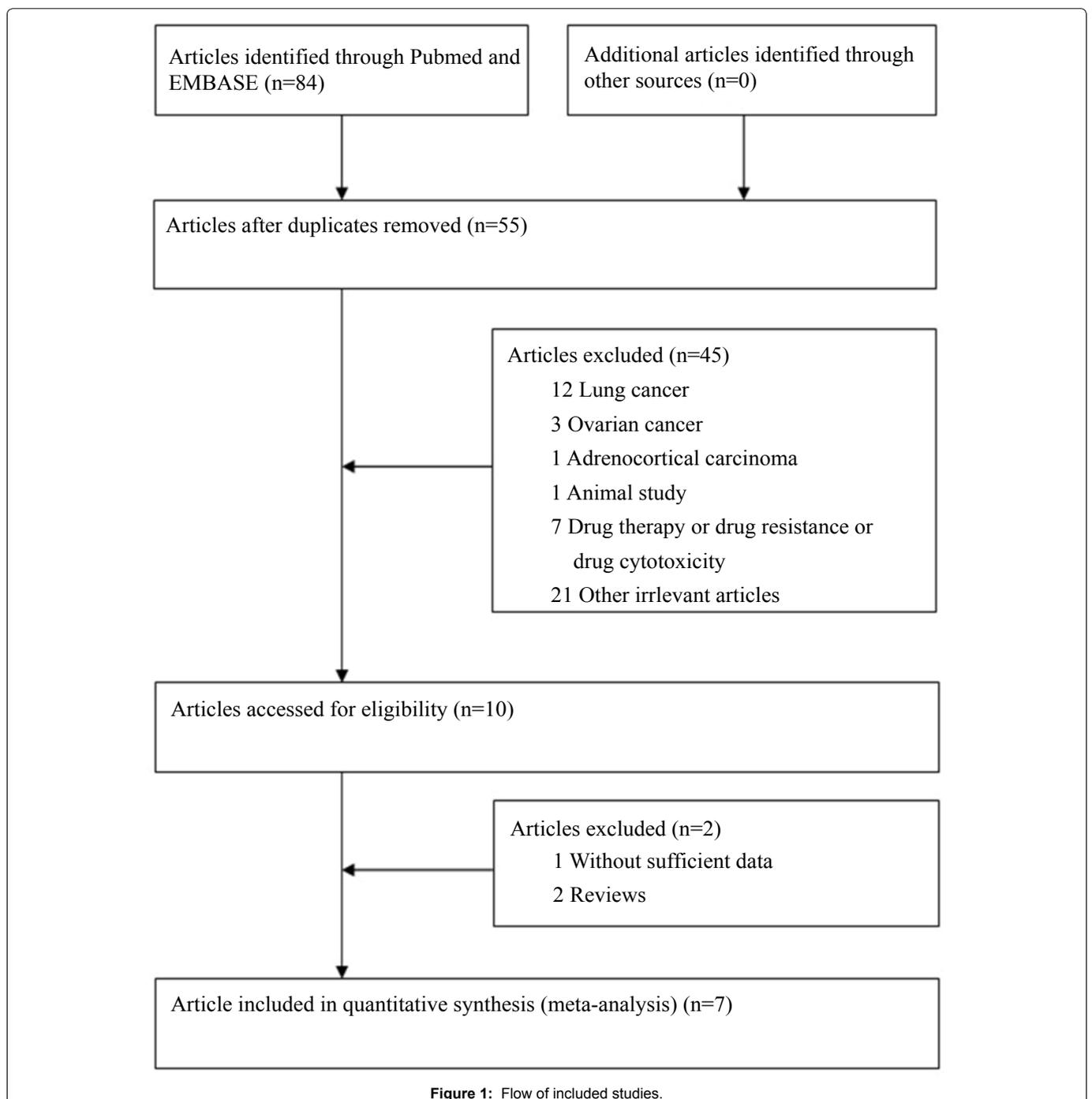
Author	Country	Ethnicity	Study design	Genotyping method	ERCC1 C8092A								ERCC1 C118T							
					No. Of case/controls	case			control			HWE	No. Of case/controls	case			control			HWE
						AA	AC	CC	AA	AC	CC			TT	CT	CC	TT	CT	CC	
Luo 2013	China	Chinese	HB	MassARRAY	-	-	-	-	-	-	-	Y	297/415	45	114	138	69	158	188	Y
Pan 2013	China	Chinese	HB	MassARRAY	443/444	45	169	229	41	162	241	Y	443/443	193	171	79	211	162	70	Y
Chen 2012	China	Chinese	HB	TaqMan	393/410	50	141	202	35	154	221	Y	393/410	68	154	171	62	152	196	Y
Zhang 2012	China	Chinese	HB	TaqMan	257/278	36	98	123	29	105	144	Y	257/278	33	94	130	22	107	149	Y
McKean-Cowdin 2009	USA	Caucasian	Mixed	Mixed	977/1870	59	361	557	55	728	1087	Y	-	-	-	-	-	-	-	Y
Wrench 2005	USA	Caucasian	PB	Mixed	450/508	36	176	238	29	212	267	Y	-	-	-	-	-	-	-	Y
Chen 2000	USA	Caucasian	PB	PCR-SSCP	122/159	6	43	73	8	70	81	Y	-	-	-	-	-	-	-	Y

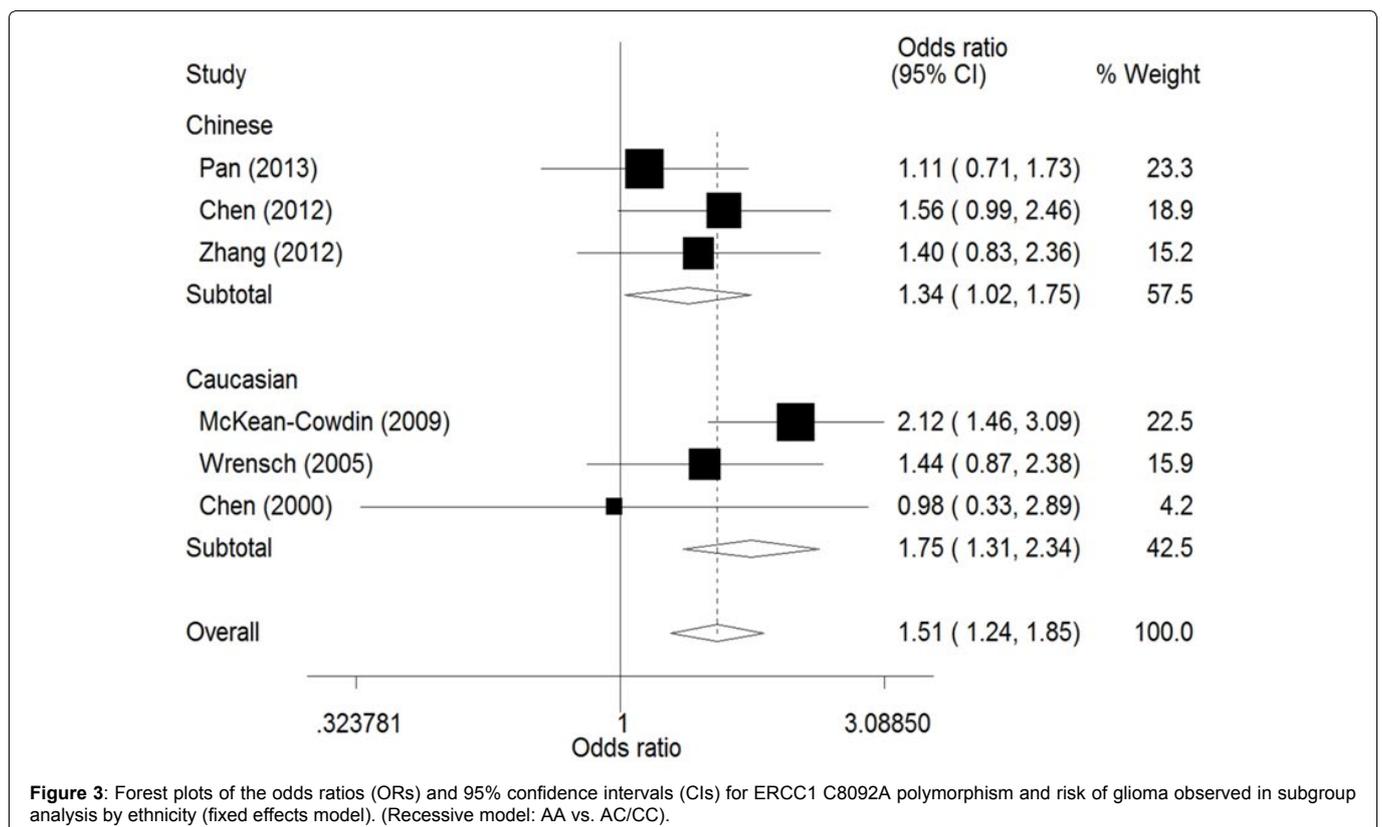
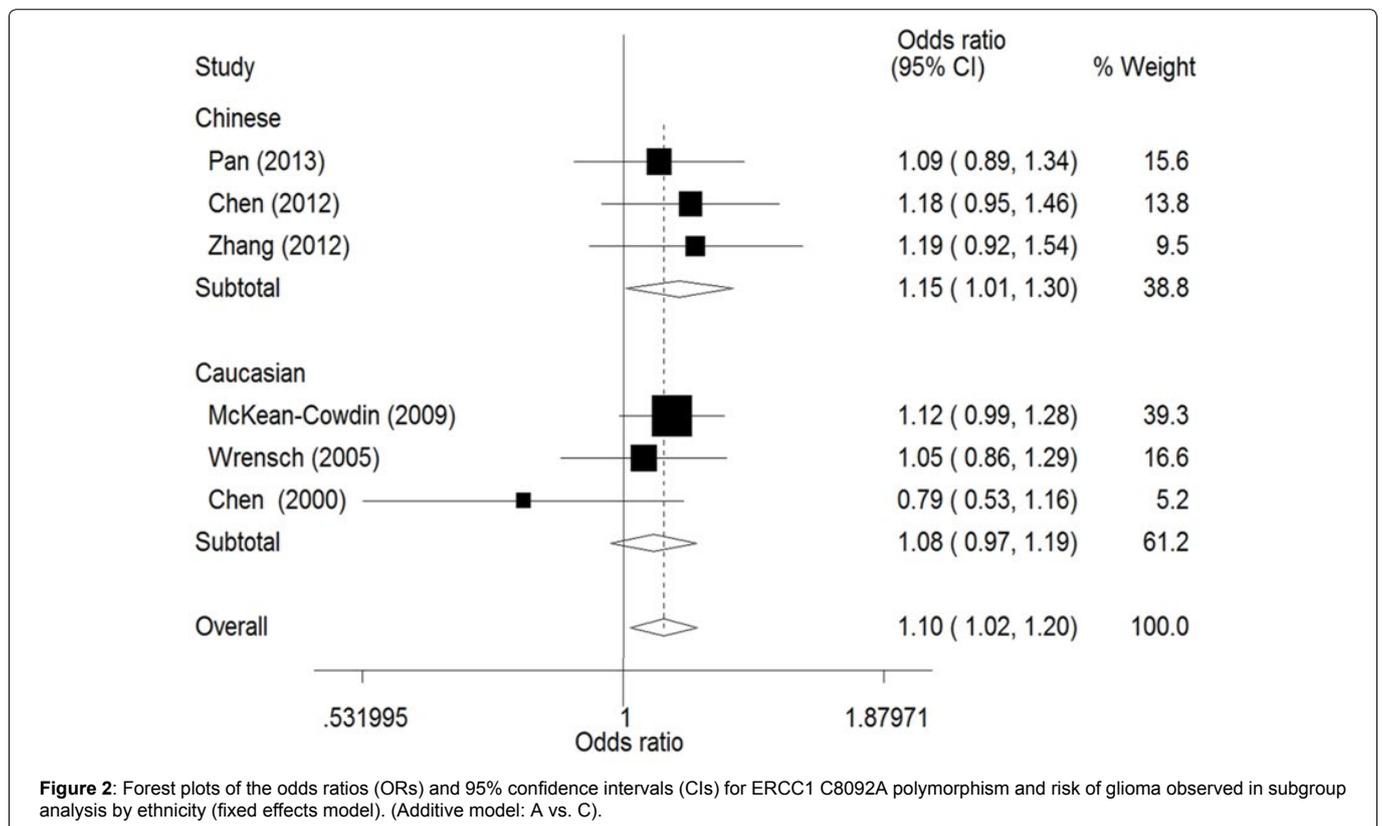
**Note.** Abbreviations: HWE, Hardy-Weinberg equilibrium; HB, hospital-based; PB, population-based; TaqMan, real-time TaqMan analysis; MassARRAY, genotyping was performed using MassARRAY platform; SSCP, Single-Strand conformation polymorphism; Y: In agreement with HWE.

**Table 2:** Results of meta-analysis for ERCC1 C8092A and C118T and the risk of glioma.

Genetic model	Dominant model			Recessive model			Co-dominant model			Co-dominant model			Additive model		
	AA/AC vs. CC			AA vs. AC/CC			AA vs. CC			AC vs. CC			A vs. C		
	OR (95% CI)	$P_h$	$I^2$ (%)												
C8092A															
Total	1.04 (0.94,1.16)	0.586	0	1.51 (1.24,1.85)	0.335	12.6	1.52 (1.24,1.86)	0.37	6.8	0.98 (0.88,1.09)	0.666	0	1.1 (1.02,1.20)	0.558	0
Ethnicity															
Chinese	1.12 (0.95,1.33)	0.962	0	1.34 (1.02,1.75)	0.566	0	1.37 (1.03,1.81)	0.65	0	1.06 (0.89,1.27)	0.894	0	1.15 (1.01,1.30)	0.832	0
Caucasian	1.00 (0.88,1.14)	0.281	21.1	1.75 (1.31,2.34)	0.259	25.9	1.70 (1.27,2.29)	0.19	39.7	0.93 (0.82,1.07)	0.418	0	1.08 (0.97,1.19)	0.22	33.9
C118T															
	TT/CT vs. CC			TT vs. CT/CC			TT vs. CC			CT vs. CC			T vs. C		
	OR (95% CI)	$P_h$	$I^2$ (%)												
Total	1.04 (0.89,1.22)	0.478	0	0.99 (0.83,1.19)	0.116	49.2	1.04 (0.84,1.29)	0.12	48.7	1.03 (0.87,1.22)	0.818	0	1.02 (0.91,1.13)	0.119	48.7

**Abbreviations:**  $P_h$ : p values for heterogeneity from Q test, random-effects model was used when p values for heterogeneity test < 0.10; otherwise, fixed-effects model was used.



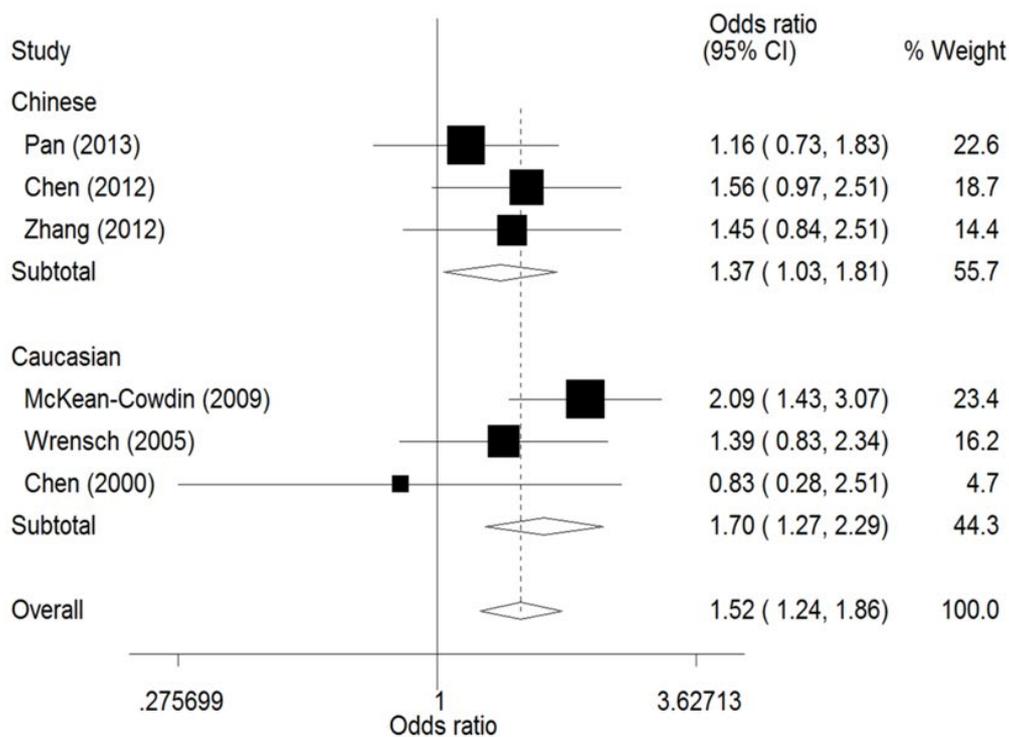


models (AA/AC vs. CC: OR = 1.04, 95% CI 0.94 - 1.06; AC vs. CC: OR = 0.98, 95% CI 0.88 - 1.09). Whereas there was significant association between ERCC1 C8092A polymorphism and susceptibility to glioma, which could be identified in these genetic models (A vs. C: OR = 1.10, 95% CI 1.02 - 1.20, [figure 2](#); AA vs. AC/CC: OR = 1.51, 95% CI 1.24-1.85, [figure 3](#), AA vs. CC: OR = 1.52, 95% CI 1.24-1.86, [figure 4](#)).

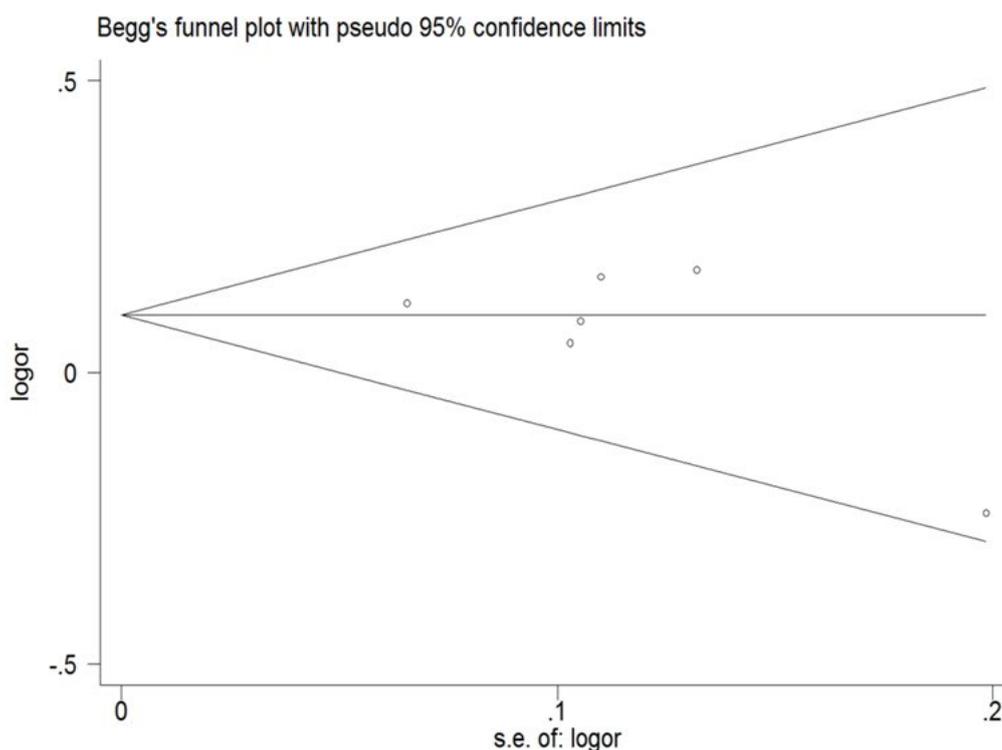
In the ethnicity subgroup analysis, as for Chinese population there was significant association between ERCC1 C8092A polymorphism and susceptibility to glioma in these genetic models (A vs. C: OR =

1.15, 95% CI 1.01 - 1.30, [figure 2](#); AA vs. AC/CC: OR = 1.34, 95% CI 1.02 - 1.75, [figure 3](#); AA vs. CC: OR = 1.37, 95% CI 1.03 - 1.81, [figure 4](#)). And in Caucasian population these genetic models were significant (AA vs. AC/CC: OR = 1.75, 95% CI 1.31 - 2.34, [figure 3](#); AA vs. CC: OR = 1.70, 95% CI 1.27 - 2.29, [figure 4](#)).

For ERCC1 C118T polymorphism, no heterogeneity was found among all of genetic models either. There was no significant association between ERCC1 C118T polymorphism and glioma risk, which could be identified in any of the genetic models (T vs. C:



**Figure 4:** Forest plots of the odds ratios (ORs) and 95% confidence intervals (CIs) for ERCC1 C8092A polymorphism and risk of glioma observed in subgroup analysis by ethnicity (fixed effects model). (Co-dominant model: AA vs. CC).



**Figure 5:** Begg's funnel plot of publication bias for ERCC1 C8092A and C118T. (Begg's funnel plot for C8092A: A vs. C). Each point represents a separate study for the indicated association; horizontal line represents the meta-analysis summary estimate; log OR, natural logarithm of OR; s.e. of logOR, standard of the log OR.

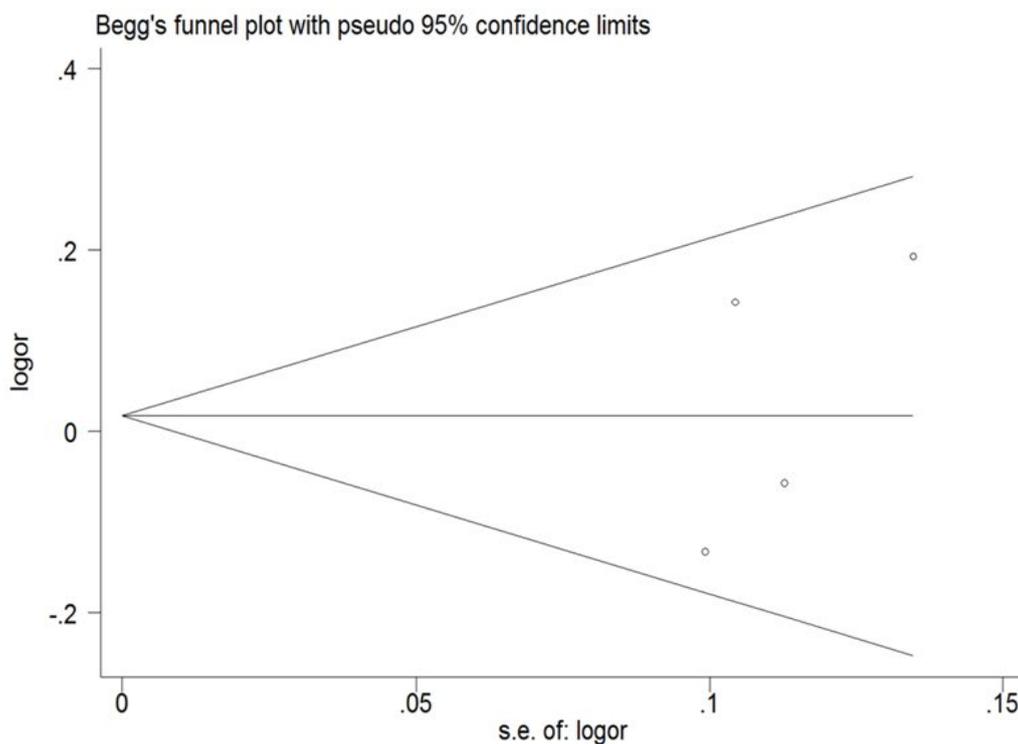
OR = 1.02, 95% CI 0.91 - 1.13; TT/CT vs. CC: OR = 1.04, 95% CI 0.89 - 1.22; TT vs. CT/CC: OR = 0.99, 95% CI 0.83 - 1.19; TT vs. CC: OR = 1.04, 95% CI 0.84 - 1.29; CT vs. CC: OR = 1.03, 95% CI 0.87 - 1.22).

### Sensitivity analysis

We deleted one single study from the overall pooled analysis each time to check the influence of the removed data set to the pooled ORs. No study was observed to change the homogeneity in heterozygote comparison.

### Bias diagnostics

Funnel plot and Egger's test were performed to assess the publication bias of the literature (Figure 5 and Figure 6). Symmetrical funnel plots were obtained in ERCC1 C8092A and C118T tested in all of the models. Egger's test further confirmed the absence of publication bias in this meta-analysis ( $P > 0.05$ ). No evidence of publication bias was observed in any comparison model.



**Figure 6:** Begg's funnel plot of publication bias for ERCC1 C8092A and C118T. (Begg's funnel plot for C118T: T vs. C). Each point represents a separate study for the indicated association; horizontal line represents the meta-analysis summary estimate; log OR, natural logarithm of OR; s.e. of logOR, standard of the log OR.

## Discussion

The NER is a highly conserved DNA repair pathway that repairs DNA lesions which alter the helical structure of the DNA molecule and interfere with DNA replication and transcription [33]. The NER pathway repairs bulky lesions such as pyrimidine dimers, other photo-products, larger chemical adducts, and cross-links. The NER pathway involves at least four steps: (a) damage recognition by a complex of bound proteins including XPC, (b) unwinding of the DNA by the TFIIH complex that includes XPD, (c) removal of the damaged single-stranded fragment (usually about 27 - 30 bp) by molecules including an ERCC1 and XPF complex, and (d) synthesis by DNA polymerases [8]. The NER pathway removes damaged DNA bases by introducing nicks 5' and 3' to an abasic site in vitro, and NER contributes to the release of 8-oxoguanine from DNA [34,35]. Variation in efficiency of these processes might influence either cancer development.

In recent years, interest in the genetic susceptibility to cancers has led to a growing attention to the study of polymorphisms of genes involved in tumorigenesis. The DNA repair gene, excision repair cross-complementing rodent repair deficiency complementation group1 (ERCC1), is potentially relevant to cancer because of their involvement in the process of nucleotide excision repair (NER) [8]. The ERCC1 gene has been widely studied, and the ERCC1 C8091A TT genotype was observed to increase the risk of lung cancer [36]. There have been a few studies reported the relationship between the ERCC1 C8092A polymorphism and glioma susceptibility [19-22,30-32]. And only a few of these studies showed the ERCC1 C8092A polymorphism was associated with the risk of glioma [21], however, some not [19,20,22,30-32]. In the present meta-analysis, the combined results based on all studies showed that ERCC1 C8092A polymorphism was significantly associated risk (additive model: OR = 1.10, 95% CI 1.02 - 1.20; recessive model: OR = 1.51, 95% CI 1.24 - 1.85; AA vs. CC: OR = 1.52, 95% CI 1.24 - 1.86). The results indicated that individuals carrying at least one T allele might be at increased risk of glioma, which in agreement with the previous study, which suggested that the T allele may serve as a dangerous biomarker [21]. Using data collected through 1994, Chen found a statistically significant association with oligoastrocytomas (OR = 4.6; 95% CI, 1.6 - 13.2), but not with other types of glioma [21]. According to described

above, our finding suggested that ERCC1 C8092A polymorphism may play a finite role on glioma risk. More studies and larger samples will be needed to further identify this relationship.

Although many studies have been performed to explore the etiology of glioma, the etiology is still not completely understood. Glioma, as a complex disease, is considered as a result of combined effects of multi-factors, including the inherited and environmental factors. Unfortunately, there are few studies focused on the gene-environment and gene-gene interactions with glioma risk. A significant and protective effect was found by Roberta when three single-nucleotide polymorphisms (ERCC2 rs13181, ERCC1 rs3212986, and GLTSCR1 rs1035938) located near each other on chromosome 19 were modeled as a haplotype. The most common haplotype (AGC) was associated with a 23% reduction in risk ( $P = 0.03$ ) compared with all other haplotypes combined [22]. So, further studies should include the gene and environmental factors and detect the potential interactions between ERCC1 C8092A polymorphism and these factors.

In our study, we did not find any relation between the risk of glioma and ERCC1 C118T polymorphism, may either due to the minor effect between ERCC1 C118T polymorphism and glioma risk or to the relatively small sample size, and that is consistent with many previous studies [29-32].

To the best of our knowledge, this is the first meta-analysis evaluating the potential association between two common polymorphisms ERCC1 C8092A and C118T and the susceptibility to glioma. However, some limitations of this meta-analysis should be acknowledged. Firstly, our finding conflicts with some articles which indicated that there was no statistical evidence to assume a correlation between ERCC1 C8092A polymorphism and glioma risk. The reason is that one of the six studies [22] included in our meta-analysis related to ERCC1 C8092A played an important role in the pooled OR. According to the data provided in the study, the OR was calculated (recessive model: AC vs. AC/CC, OR = 2.09, 95% CI 1.43 - 3.07), that indicated the ERCC1 C8092A TT genotype increased glioma risk. However, after adjusted for age, gender, and study center, the relationship was not significant. The meta-regression method should be used to adjust the confounders in future study. Secondly, the target population of one study included in this meta-analysis was mainly on Caucasian population [20], but there was not enough data about

Caucasian population provided by this paper. In our study, we used the information of all the population instead of the Caucasian population.

Despite some limitations, this meta-analysis indicated that subjects carrying at least one A allele of ERCC1 C8092A might increase the risk of glioma, whereas ERCC1 C118T polymorphism might have no influence on the susceptibility of glioma.

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## Conflict of Interest

The authors declare that they have no conflict of interest.

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